STT 3850 : Week 5

Spring 2024

Appalachian State University

Section 1

Outline for the week

- Simple Linear Regression (One categorical explanatory variable)
- Multiple Regression (One numerical and one categorical explanatory variable)
- Interaction model
- Two numerical explanatory variables

Section 2

One categorical explanatory variable

- It's an unfortunate truth that life expectancy is not the same across all countries in the world.
 - International development agencies are interested in studying these differences in life expectancy in the hopes of identifying where governments should allocate resources to address this problem.

• In this section, we'll explore differences in life expectancy in two ways:

- Differences between continents: Are there significant differences in average life expectancy between the five populated continents of the world: Africa, the Americas, Asia, Europe, and Oceania?
- Differences within continents: How does life expectancy vary within the world's five continents? For example, is the spread of life expectancy among the countries of Africa larger than the spread of life expectancy among the countries of Asia?

- To answer such questions, we'll use the gapminder data frame included in the gapminder package
 - This dataset has international development statistics such as life expectancy, GDP per capita, and population for 142 countries for 5-year intervals between 1952 and 2007.
 - Recall we visualized some of this data earlier.
- We'll use this data for basic regression again, but now using an explanatory variable x that is **categorical**.
 - $\bullet\,$ A numerical outcome variable y (a country's life expectancy) and
 - A single categorical explanatory variable x (the continent that the country is a part of).

Loading needed packages.

library(tidyverse)
library(moderndive)
library(skimr)

library(gapminder)

loading collection of packages
datasets and regression functions
provides a simple-to-use functions
for summary statistics
datasets

- The data on the 142 countries can be found in the gapminder data frame included in the gapminder package.
 - However, to keep things simple, let's filter() for only those observations/rows corresponding to the year 2007.
 - Additionally, let's select() only the subset of the variables we'll consider in this chapter. We'll save this data in a new data frame called gapminder2007:

```
gapminder2007 <- gapminder %>%
filter(year == 2007) %>%
select(country, lifeExp, continent, gdpPercap)
glimpse(gapminder2007)
```

Rows: 142 Columns: 4 \$ country <fct> "Afghanistan", "Albania", "Algeria", "Angola", "A \$ lifeExp <dbl> 43.828, 76.423, 72.301, 42.731, 75.320, 81.235, 7 \$ continent <fct> Asia, Europe, Africa, Africa, Americas, Oceania, 1 \$ gdpPercap <dbl> 974.5803, 5937.0295, 6223.3675, 4797.2313, 12779.3

- A full description of all the variables included in gapminder can be found by reading the associated help file (run ?gapminder in the console).
- However, let's fully describe the 4 variables we selected in gapminder2007:
 - country: An identification variable of type character/text used to distinguish the 142 countries in the dataset.
 - IifeExp: A numerical variable of that country's life expectancy at birth. This is the outcome variable y of interest.
 - Continent: A categorical variable with five levels. Here "levels" correspond to the possible categories: Africa, Asia, Americas, Europe, and Oceania. This is the explanatory variable x of interest.
 - gdpPercap: A numerical variable of that country's GDP per capita in US inflation-adjusted dollars.

Let's look at a random sample of five out of the 142 countries

```
gapminder2007 %>%
sample_n(size = 5)
```

```
# A tibble: 5 x 4
```

	country	lifeExp	continent	gdpPercap
	<fct></fct>	<dbl></dbl>	<fct></fct>	<dbl></dbl>
1	Cote d'Ivoire	48.3	Africa	1545.
2	Venezuela	73.7	Americas	11416.
3	Nigeria	46.9	Africa	2014.
4	South Africa	49.3	Africa	9270.
5	Lesotho	42.6	Africa	1569.

Let's compute the summary statistics using the skim() function

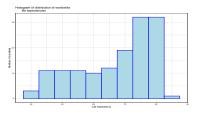
```
gapminder2007 %>%
  select(lifeExp, continent) %>%
  skim()
```

```
# Or using summary()
gapminder2007 %>%
select(lifeExp, continent) %>%
summary()
```

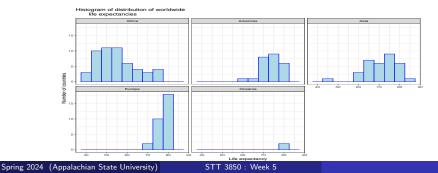
life	Exp	continent			
Min.	:39.61	Africa	:52		
1st Qu.	:57.16	Americas	:25		
Median	:71.94	Asia	:33		
Mean	:67.01	Europe	:30		
3rd Qu.	:76.41	Oceania	: 2		
Max.	:82.60				

Why is the mean life expectancy lower than the median?

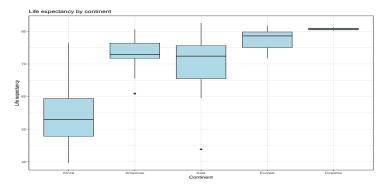
```
ggplot(gapminder2007, aes(x = lifeExp)) +
  geom_histogram(binwidth=5, color = "blue", fill = "lightblue") +
  labs(x = "Life expectancy", y = "Number of countries",
        title = "Histogram of distribution of worldwide
        life expectancies") +
  theme_bw()
```



We see that this data is skewed to the left \rightarrow mean < median.



```
ggplot(gapminder2007, aes(x = continent, y = lifeExp)) +
geom_boxplot(fill = "lightblue") +
labs(x = "Continent", y = "Life expectancy",
        title = "Life expectancy by continent") +
theme_bw()
```



continent	median	mean	Difference versus Africa
Africa	52.9265	54.80604	0.00000
Americas	72.8990	73.60812	18.80208
Asia	72.3960	70.72848	15.92245
Europe	78.6085	77.64860	22.84256
Oceania	80.7195	80.71950	25.91346

In our life expectancy example, we now instead have a **categorical** explanatory variable continent:

```
lifeExp_model <- lm(lifeExp ~ continent, data = gapminder2007)
knitr::kable(get_regression_table(lifeExp_model))</pre>
```

term	estimate	std_error	statistic	p_value	lower_ci	upper_ci
intercept	54.806	1.025	53.446	0	52.778	56.834
continent: Americas	18.802	1.800	10.448	0	15.243	22.361
continent: Asia	15.922	1.646	9.675	0	12.668	19.177
continent: Europe	22.843	1.695	13.474	0	19.490	26.195
continent: Oceania	25.913	5.328	4.863	0	15.377	36.450

Our model will not yield a "best-fitting" regression line like when the x is continuous, but rather offsets relative to a baseline for comparison.

Let's break the 5 estimates down one-by-one:

- intercept corresponds to the mean life expectancy of countries in Africa of 54.8 years.
- continent: Americas corresponds to countries in the Americas and the value +18.8 is the same difference in mean life expectancy relative to Africa we displayed earlier. In other words, the mean life expectancy of countries in the Americas is 54.8 + 18.8 = 73.6.
- continent: Asia the mean life expectancy of countries in Asia is 54.8 + 15.9 = 70.7.
- continent: Europe the mean life expectancy of countries in Europe is 54.8 + 22.8 = 77.6.
- continent: Oceania the mean life expectancy of countries in Oceania is 54.8 + 25.9 = 80.7.

We can change the baseline group to be another continent. In what follows the baseline is changed to be Americas instead of Africa.

gapminder2007\$continent <- relevel(gapminder2007\$continent, ref='Americas')
lifeExp_model1 <- lm(lifeExp - continent, data = gapminder2007)
knitr::kable(get_regression_table(lifeExp_model1))</pre>

term	estimate	std_error	statistic	p_value	lower_ci	upper_ci
intercept	73.608	1.479	49.772	0.000	70.684	76.533
continent: Africa	-18.802	1.800	-10.448	0.000	-22.361	-15.243
continent: Asia	-2.880	1.961	-1.469	0.144	-6.757	0.997
continent: Europe	4.040	2.002	2.018	0.046	0.081	8.000
continent: Oceania	7.111	5.434	1.309	0.193	-3.634	17.857

The equation for our fitted values for model (lifeExp_model) is written as:

$$\begin{split} \hat{y} &= \widehat{\mathsf{life} \, \exp} = b_0 + b_{\mathsf{Amer}} \cdot \mathbf{1}_{\mathsf{Amer}}(x) + b_{\mathsf{Asia}} \cdot \mathbf{1}_{\mathsf{Asia}}(x) \\ &+ b_{\mathsf{Euro}} \cdot \mathbf{1}_{\mathsf{Euro}}(x) + b_{\mathsf{Ocean}} \cdot \mathbf{1}_{\mathsf{Ocean}}(x) \\ &= 54.8 + 18.8 \cdot \mathbf{1}_{\mathsf{Amer}}(x) + 15.9 \cdot \mathbf{1}_{\mathsf{Asia}}(x) \\ &+ 22.8 \cdot \mathbf{1}_{\mathsf{Euro}}(x) + 25.9 \cdot \mathbf{1}_{\mathsf{Ocean}}(x). \end{split}$$

where for example:

$$1_{\mathsf{Amer}}(x) = \begin{cases} 1 & \text{if country } x \text{ is in the Americas} \\ 0 & \text{if otherwise} \end{cases}$$

Let's put this all together and compute the fitted value $\hat{y} = |\widetilde{\text{ife} \exp}|$ for a country in Africa.

• Since the country is in Africa, all four indicator functions $1_{\mathsf{Euro}}(x) = 1_{\mathsf{Amer}}(x) = 1_{\mathsf{Asia}}(x) = 1_{\mathsf{Ocean}}(x) = 0.$

$$\begin{split} \hat{y} &= \widehat{\mathsf{life exp}} &= 54.8 + 18.8 \cdot 1_{\mathsf{Amer}}(x) + 15.9 \cdot 1_{\mathsf{Asia}}(x) \\ &+ 22.8 \cdot 1_{\mathsf{Euro}}(x) + 25.9 \cdot 1_{\mathsf{Ocean}}(x). \\ &= 54.8 + 18.8 \cdot 0 + 15.9 \cdot 0 + 22.8 \cdot 0 + 25.9 \cdot 0 \\ &= 54.8 \end{split}$$

For a country in in the Americas, in this case, only the indicator function $1_{Amer}(x) = 1$.

$$\hat{y} = \widehat{\text{life exp}} = 54.8 + 18.8 \cdot 1_{\text{Amer}}(x) + 15.9 \cdot 1_{\text{Asia}}(x) +22.8 \cdot 1_{\text{Euro}}(x) + 25.9 \cdot 1_{\text{Ocean}}(x). = 54.8 + 18.8 \cdot 1 + 15.9 \cdot 0 + 22.8 \cdot 0 + 25.9 \cdot 0 = 73.6$$

In general, if we fit a linear regression model using a categorical explanatory variable x that has k possible categories, the regression table will return an intercept and k-1 offsets.

Observed/fitted values and residuals

knitr::kable(regression_points %>% head(n = 9))

country	lifeExp	continent	lifeExp_hat	residual
Afghanistan	43.828	Asia	70.728	-26.900
Albania	76.423	Europe	77.649	-1.226
Algeria	72.301	Africa	54.806	17.495
Angola	42.731	Africa	54.806	-12.075
Argentina	75.320	Americas	73.608	1.712
Australia	81.235	Oceania	80.720	0.516
Austria	79.829	Europe	77.649	2.180
Bahrain	75.635	Asia	70.728	4.907
Bangladesh	64.062	Asia	70.728	-6.666

Section 3

Multiple Regression

In the previous chapter, we introduced ideas related to modeling for **explanation.**

- In particular that the goal of modeling is to make explicit the relationship between some outcome variable y and some explanatory variable x.
- We focused on linear regression, where we only considered one explanatory x variable that is either numeric or categorical.
- Now, we we'll start considering models that include more than one explanatory variable *x*.
 - NOTE: the interpretation of the associated effect of any one explanatory variable must be made in conjunction with the other explanatory variables included in your model.

Let's load all the packages needed for this chapter.

library(tidyverse)
library(moderndive)
library(skimr)
library(ISLR)

One numerical and one categorical explanatory variable

Let's revisit the instructor evaluation data we introduced earlier.

- We studied the relationship between teaching evaluation scores as given by students and "beauty" scores.
 - The variable teaching score was the numerical outcome variable y, and the variable "beauty" score (bty_avg) was the numerical explanatory x variable.
- In this section, we are going to consider a different model
 - Our outcome variable will still be teaching score, but
 - we'll now include two different explanatory variables: age and gender.
 - Could it be that instructors who are older receive better teaching evaluations from students?
 - Or could it instead be that younger instructors receive better evaluations?
 - Are there differences in evaluations given by students for instructors of different genders?

Let's select() only the subset of the variables we'll consider in this chapter.

evals_ch6 <- evals %>%
 select(ID, score, age, gender)

Recall the three common steps in an exploratory data analysis:

- Looking at the raw data values.
- Omputing summary statistics.
- Oreating data visualizations.

Let's first look at the raw data values.

```
glimpse(evals_ch6)
```

```
Rows: 463
Columns: 4
$ ID <int> 1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 14, 15, 1
$ score <dbl> 4.7, 4.1, 3.9, 4.8, 4.6, 4.3, 2.8, 4.1, 3.4, 4.5, 3.4
$ gender <fct> female, female, female, female, male, male, male, male
# Or
evals ch6 %>%
 sample n(size = 2)
# A tibble: 2 \times 4
    ID score age gender
 <int> <dbl> <int> <fct>
1
 434 2.8 62 male
2
   208 4.4 62 male
```

```
evals_ch6 %>%
   select(score, age, gender) %>%
   skim()
```

Or

```
evals_ch6 %>%
   select(score, age, gender) %>%
   summary()
```

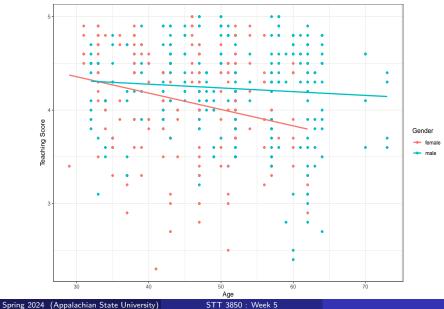
score		ag	ge	gender	
Min.	:2.300	Min.	:29.00	female	e:195
1st Qu.	:3.800	1st Qu	.:42.00	male	:268
Median	:4.300	Median	:48.00		
Mean	:4.175	Mean	:48.37		
3rd Qu.	:4.600	3rd Qu	.:57.00		
Max.	:5.000	Max.	:73.00		

Let's compute the correlation coefficient between our two numerical variables: score and age:

```
evals_ch6 %>%
  summarize(r = cor(score, age))
# A tibble: 1 x 1
       r
   <dbl>
1 - 0.107
# or using the get_correlation wrapper
# from moderndive
evals ch6 %>%
  get correlation(score ~ age)
# A tibble: 1 \times 1
     cor
   <dbl>
1 - 0.107
```

```
ggplot(evals_ch6, aes(x = age, y = score, color = gender)) +
geom_point() +
labs(x = "Age", y = "Teaching Score", color = "Gender",
    title = "Interaction Model") +
geom_smooth(method = "lm", se = FALSE) +
theme_bw() -> int_mod
int_mod
```

Interaction Model



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- Let's now quantify the relationship of our outcome variable y and the two explanatory variables using one type of multiple regression model known as an **interaction model**.
- Going back to our multiple regression model for teaching score using age and gender in the figure above, we generate the regression table using the same two-step approach.
 - First, "fit" a model using the lm() (linear model) function of the form y ~x1 + x2 + x1:x2 which is the same as y ~ x1*x2 in R's modeling notation.
 - Second, apply get_regression_table() or summary() to the linear model object created in 1.

Get regression table: knitr::kable(get_regression_table(score_model_interaction))

term	estimate	std_error	statistic	p_value	lower_ci	upper_ci
intercept	4.883	0.205	23.795	0.000	4.480	5.286
age	-0.018	0.004	-3.919	0.000	-0.026	-0.009
gender: male	-0.446	0.265	-1.681	0.094	-0.968	0.076
age:gendermale	0.014	0.006	2.446	0.015	0.003	0.024

Female instructors are the "baseline for comparison" group.

- The estimate for intercept is the intercept for only the female instructors.
- The estimate for age is the slope for age for only the female instructors.
- Thus, the red regression line in Figure 1 has an intercept of 4.883 and slope for age of -0.018.
- Note: The intercept has no practical interpretation since instructors can not have a **zero** age.

What about the intercept and slope for age of the male instructors in the blue line of Figure 1?

- The value for gender: male that appears in the Table (-0.446) is not the intercept for the male instructors but rather the offset in intercept for male instructors relative to female instructors.
 - The intercept for the male instructors is intercept + gender:male = 4.883 + (-0.446) = 4.883 0.446 = 4.437.
- Similarly, age:gendermale = 0.014 is not the slope for age for the male instructors, but rather the offset in slope for the male instructors.
 - Therefore, the slope for age for the male instructors is age + age:gendermale =-0.018 + 0.014 = -0.004.

Gender	Intercept	Slope for age
Female instructors	4.883	-0.018
Male instructors	4.437	-0.004

- Since the slope for age for the female instructors was -0.018, it means that on average, a female instructor who is a year older would have a teaching score that is 0.018 units **lower**.
- For the male instructors, the associated decrease in score is 0.004 units.
- While both slopes for age were negative, the slope for age for the female instructors is **larger** in magnitude.

Let's now write the equation for our regression lines, which we can use to compute our fitted values

$$\hat{y} = \widehat{\text{score}} = b_0 + b_{\text{age}} \cdot \text{age} + b_{\text{male}} \cdot 1_{\text{is male}}(x) + b_{\text{age:gender}} \cdot \text{age} \cdot 1_{\text{is male}}(x)$$

= 4.883 - 0.018 \cdot age - 0.446 \cdot 1_{\text{is male}}(x) + 0.014 \cdot age \cdot 1_{\text{is male}}(x).

where:

$$1_{\text{is male}}(x) = \left\{ \begin{array}{ll} 1 & \text{if instructor } x \text{ is male} \\ 0 & \text{otherwise} \end{array} \right.$$

Let's put this all together and compute the fitted value $\hat{y} = \widehat{\text{score}}$ for female instructors.

• Since for female instructors $1_{is male}(x) = 0$.

$$\hat{y} = \widehat{\text{score}} = 4.883 - 0.018 \cdot \text{age} - 0.446 \cdot 0 + 0.014 \cdot \text{age} \cdot 0.$$

= $4.883 - 0.018 \cdot \text{age}$

• For male instructors $1_{is male}(x) = 1$.

$$\begin{split} \hat{y} &= \widehat{\mathsf{score}} &= 4.883 - 0.018 \cdot \mathsf{age} - 0.446 \cdot 1 + 0.014 \cdot \mathsf{age} \cdot 1. \\ &= (4.883 - 0.446) + (-0.018 + 0.014) \cdot \mathsf{age} \\ &= 4.437 - 0.004 \cdot \mathsf{age} \end{split}$$

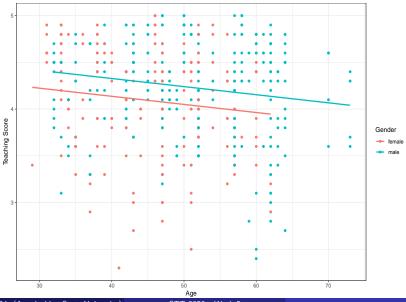
- The term $b_{\text{age:gender}}$ in the equation for the fitted value $\hat{y} = \widehat{\text{score}}$ is what's known in statistical modeling as an **interaction effect**.
- We say there is an interaction effect if the associated effect of one variable depends on the value of another variable.
 - Here, the associated effect of the variable age depends on the value of the other variable gender.
 - The difference in slopes for age of +0.014 of male instructors relative to female instructors shows this.

With one numerical and one categorical explanatory variable, another type of model we can use is known as a **parallel slopes** model.

• Unlike interaction models, parallel slopes models still allow for different intercepts but force all lines to have the same slope.

Parallel slopes model

Parallel Slopes Model



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Parallel slopes model

Get regression table:

knitr::kable(get_regression_table(score_model_parallel_slopes))

term	estimate	std_error	statistic	p_value	lower_ci	upper_ci
intercept	4.484	0.125	35.792	0.000	4.238	4.730
age	-0.009	0.003	-3.280	0.001	-0.014	-0.003
gender: male	0.191	0.052	3.632	0.000	0.087	0.294

Gender	Intercept	Slope for age
Female instructors	4.484	-0.009
Male instructors	4.675	-0.009

Let's now write the equation for our regression lines, which we can use to compute our fitted values

$$\begin{split} \hat{y} &= \widehat{\text{score}} \quad = b_0 + b_{\text{age}} \cdot \text{age} + b_{\text{male}} \cdot 1_{\text{is male}}(x) \\ &= 4.484 - 0.009 \cdot \text{age} + 0.191 \cdot 1_{\text{is male}}(x). \end{split}$$

Let's put this all together and compute the fitted value $\hat{y} = \widehat{\text{score}}$ for female instructors.

• Since for female instructors $1_{is male}(x) = 0$.

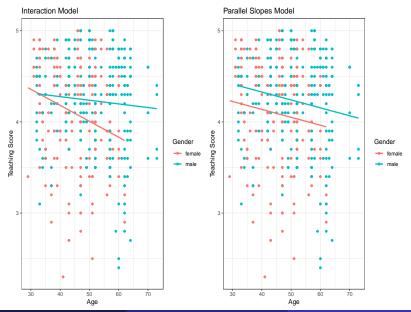
$$\hat{y} = \widehat{\text{score}} = 4.484 - 0.009 \cdot \text{age} + 0.191 \cdot 0$$

= 4.484 - 0.009 \cdot age

• For male instructors $1_{is male}(x) = 1$.

$$\begin{split} \hat{y} &= \widehat{\text{score}} &= 4.484 - 0.009 \cdot \text{age} + 0.191 \cdot 1 \\ &= (4.484 + 0.191) - (0.009) \cdot \text{age} \\ &= 4.675 - 0.009 \cdot \text{age} \end{split}$$

Interaction Model and Parallel Slopes Model

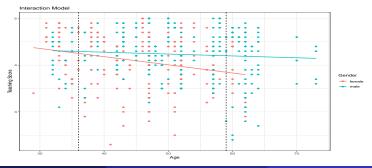


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Observed/fitted values and residuals

- We'll compute the observed values, fitted values, and residuals for the interaction model which we saved in score_model_interaction.
 - Say, you have an instructor who identifies as female and is 36 years old. What fitted value $\hat{y} = \widehat{\text{score}}$ would our model yield?
 - Say, you have another instructor who identifies as male and is 59 years old. What would their fitted value \hat{y} be?
- See if you can answer this question visually.



• For female instructors, we have.

$$\hat{y} = \widehat{\text{score}} = 4.883 - 0.018 \cdot \text{age}$$

= 4.883 - 0.018 \cdot 36 = 4.24

• For male instructors.

$$\hat{y} = \hat{\text{score}} = 4.437 - 0.004 \cdot \text{age}$$

= 4.437 - 0.004 \cdot 59 = 4.20

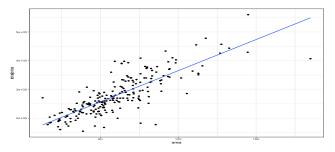
Note: It is better to let R compute the values and round at the end.

4.201373

Suppose a realtor wants to model the appraised price of an apartment as a function of the predictors living area (in m^2) and the presence or absence of elevators. Consider the data frame 'VIT2005', which contains data about apartments in Vitoria, Spain, including **totalprice**, area, and elevator, which are the appraised apartment value in Euros, living space in square meters, and the absence or presence of at least one elevator in the building, respectively. The realtor first wants to know if there is any relationship between appraised price (Y) and living area (x_1) . Next, the realtor wants to know how adding a dummy variable for whether or not an elevator is present changes the relationship: Are the lines the same? Are the slopes the same? Are the intercepts the same?

Solution (is there a realationship between totalprice and area?):

```
library(PASWR2)
VIT2005 <- VIT2005 %>%
mutate(elevator = factor(elevator, labels = c("No", "Yes")))
ggplot(data = VIT2005, aes(x = area, y = totalprice)) +
geom_point() +
theme_bw() +
geom_smooth(method = "lm", se = FALSE)
```



Solution (is there a realationship between totalprice and area?):

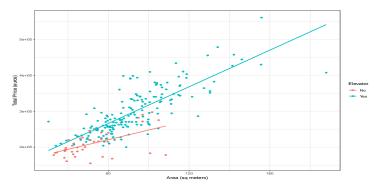
```
mod_simple <- lm(totalprice ~ area, data = VIT2005)</pre>
summary(mod_simple)
Call:
lm(formula = totalprice ~ area, data = VIT2005)
Residuals:
   Min 10 Median 30
                                  Max
-156126 -21564 -2155 19493 120674
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 40822.4 12170.1 3.354 0.00094 ***
             2704.8 133.6 20.243 < 2e-16 ***
area
___
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 40810 on 216 degrees of freedom
Multiple R-squared: 0.6548, Adjusted R-squared: 0.6532
F-statistic: 409.8 on 1 and 216 DF, p-value: < 2.2e-16
```

Solution (is there a realationship between totalprice and area?):

knitr::kable(get_regression_table(mod_simple))

term	estimate	std_error	statistic	p_value	lower_ci	upper_ci
intercept	40822.416	12170.092	3.354	0.001	16835.075	64809.757
area	2704.751	133.616	20.243	0.000	2441.393	2968.109

Solution (does adding a dummy variable (elevator) change the relationship?):



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Solution (does adding a dummy variable (elevator) change the relationship?):

mod_int <- lm(totalprice ~ area + elevator + area:elevator, data = VIT2005)</pre> summary(mod int) Call: lm(formula = totalprice ~ area + elevator + area:elevator, data = VIT2005) Residuals: Min 10 Median 30 Max -133610 -22216 -2423 20276 113159 Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 113114.0 28985.0 3.902 0.000128 *** 1343.7 392.2 3.426 0.000735 *** area elevatorYes -50871.7 31990.6 -1.590 0.113264 area:elevatorYes 1202.0 417.4 2.880 0.004380 ** ___ Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 37610 on 214 degrees of freedom

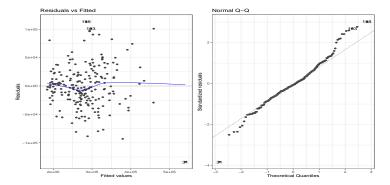
Multiple R-squared: 0.7096, Adjusted R-squared: 0.7055 F-statistic: 174.3 on 3 and 214 DF, p-value: < 2.2e-16

Solution (does adding a dummy variable (elevator) change the relationship?):

```
mod_ps <- lm(totalprice ~ area + elevator, data = VIT2005)</pre>
summary(mod_ps)
Call:
lm(formula = totalprice ~ area + elevator, data = VIT2005)
Residuals:
   Min 10 Median 30
                                  Max
-120265 -20224 -2567 18281 112406
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 36173.6 11434.8 3.163 0.00178 **
area 2405.4 136.3 17.652 < 2e-16 ***
elevatorYes 39091.1 7022.8 5.566 7.71e-08 ***
___
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 38240 on 215 degrees of freedom
Multiple R-squared: 0.6983, Adjusted R-squared: 0.6955
F-statistic: 248.8 on 2 and 215 DF, p-value: < 2.2e-16
```

Diagnostic Plots

library(ggfortify)
autoplot(mod_int, ncol = 2, nrow = 1, which = 1:2) +
 theme_bw()



Section 4

Two numerical explanatory variables

- Let's switch gears and consider multiple regression models where instead of one numerical and one categorical explanatory variable, we have two numerical explanatory variables.
- The Credit dataset we will use is from the ISLR package.
 - The outcome variable of interest is the credit card debt of 400 individuals.
 - Other variables like income, credit limit, credit rating, and age are included as well.
- Note that the Credit data is not based on real individuals' financial information, but rather is a simulated dataset used for educational purposes.

Use select() to create a subset of the variables we'll consider in this chapter.

Rows: 400

Columns: 6

\$ ID <int> 1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 14 \$ debt <int> 333, 903, 580, 964, 331, 1151, 203, 872, 279, \$ credit_limit <int> 3606, 6645, 7075, 9504, 4897, 8047, 3388, 711 \$ income <dbl> 14.891, 106.025, 104.593, 148.924, 55.882, 80 \$ credit_rating <int> 283, 483, 514, 681, 357, 569, 259, 512, 266, 4 \$ age <int> 34, 82, 71, 36, 68, 77, 37, 87, 66, 41, 30, 6

```
credit_ch6 %>%
  sample_n(size = 5)
```

```
# A tibble: 5 \ge 6
    ID debt credit_limit income credit_rating
                                             age
 <int> <int>
                   <int> <dbl>
                                      <int> <int>
1
   101 298
                   3736 21.2
                                       256
                                              41
2 236 191
                   2923 10.5
                                       232
                                              25
3
  83 503
                   4433 23.7
                                       344 63
4
 357 962
                   6090 34.5
                                       442 36
5
   115 271
                   3326 16.5
                                       268
                                              41
credit ch6 %>%
 select(debt, credit limit, income) %>%
 skim()
```

credit_ch6 %>% select(debt, credit_limit, income) %>% summary()

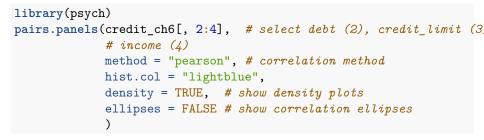
debt	credit_limit	income	
Min. : 0.00	Min. : 855	Min. : 10.35	
1st Qu.: 68.75	1st Qu.: 3088	1st Qu.: 21.01	
Median : 459.50	Median : 4622	Median : 33.12	
Mean : 520.01	Mean : 4736	Mean : 45.22	
3rd Qu.: 863.00	3rd Qu.: 5873	3rd Qu.: 57.47	
Max. :1999.00	Max. :13913	Max. :186.63	

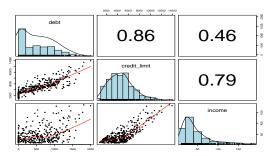
We can compute the correlation coefficient between the different possible pairs of these variables.

credit_ch6 %>%
 select(debt, credit_limit, income) %>%
 cor()

	debt	${\tt credit_limit}$	income
debt	1.0000000	0.8616973	0.4636565
credit_limit	0.8616973	1.0000000	0.7920883
income	0.4636565	0.7920883	1.0000000

Exploratory data analysis



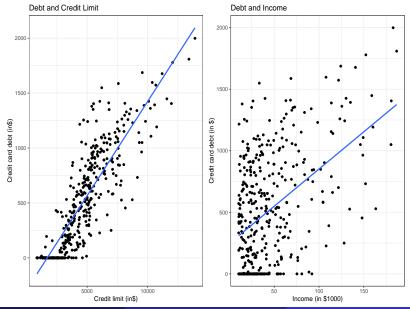


Exploratory data analysis: Collinearity

- We say there is a high degree of collinearity between the credit_limit and income explanatory variables.
- Collinearity (or multicollinearity) is a phenomenon where one explanatory variable in a multiple regression model is **highly** correlated with another.
- So in our case since credit_limit and income are highly correlated.
 - If we knew a persons' credit_limit, we could make a pretty good guess about their income.
 - Thus, these two variables provide somewhat redundant information.
- We will leave discussion on how to work with collinear explanatory variables for another course.

Let's visualize the relationship of the outcome variable with each of the two explanatory variables in two separate plots

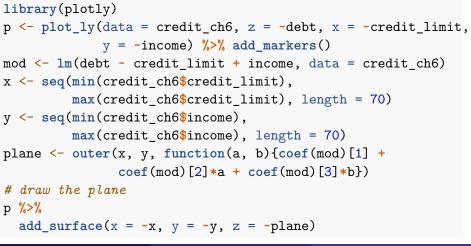
```
ggplot(data = credit_ch6, aes(x = credit_limit, y = debt)) +
  geom point() +
  labs(x= "Credit limit (in$)", y = "Credit card debt (in$)",
       title = "Debt and Credit Limit") +
  geom_smooth(method = "lm", se = FALSE) +
  theme bw() \rightarrow p1
ggplot(data = credit_ch6, aes(x = income, y = debt)) +
  geom point() +
  labs(x = "Income (in $1000)", y = "Credit card debt (in $)"
       title = "Debt and Income") +
  geom_smooth(method = "lm", se = FALSE) +
  theme_bw() \rightarrow p2
library(patchwork)
p1 + p2
```

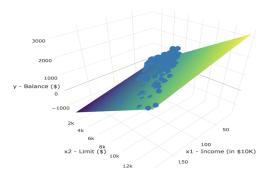


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To visualize the joint relationship of all three variables simultaneously, we need a 3-dimensional (3D) scatterplot. The following code will create a 3-dimensional scatterplot.





The regression plane is the "best-fitting" plane that similarly minimizes the sum of squared residuals.

term	estimate	std_error	statistic	p_value	lower_ci	upper_ci
intercept	-385.179	19.465	-19.789	0	-423.446	-346.912
credit_limit	0.264	0.006	44.955	0	0.253	0.276
income	-7.663	0.385	-19.901	0	-8.420	-6.906

Regression plane: Interpretation

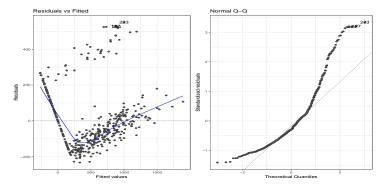
- First, the intercept value is -\$385.179.
 - This intercept represents the credit card debt for an individual who has credit_limit of \$0 and income of \$0.
 - In our data, the intercept has no practical interpretation since no individuals had both credit_limit an income values of \$0.
 - Rather, the intercept is used to situate the regression plane in 3D space.
- Second, the credit_limit value is \$0.264.
 - Taking into account all the other explanatory variables in our model, for every increase of one dollar in credit_limit, there is an associated increase of on average \$0.26 in credit card debt.
 - Just as we earlier, we are cautious not to imply causality. We do this merely stating there was an associated increase.
- Third, income = -\$7.66.
 - Taking into account all other explanatory variables in our model, for every increase of one unit of income (\$1000 in actual income), there is an associated decrease of, on average, \$7.66 in credit card debt.

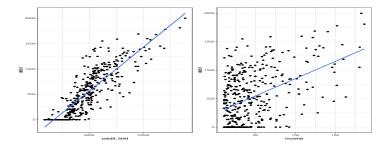
Putting these results together, the equation of the regression plane that gives us fitted values $\hat{y} = \widehat{\text{debt}}$ is:

$$\begin{aligned} \hat{y} &= b_0 + b_1 \cdot x_1 + b_2 \cdot x_2 \\ \widehat{\mathsf{debt}} &= b_0 + b_{limit} \cdot \mathsf{limit} + b_{income} \cdot \mathsf{income} \\ &= -385.179 + 0.263 \cdot \mathsf{limit} - 7.663 \cdot \mathsf{income} \end{aligned}$$

Diagnostic Plots

autoplot(debt_model, ncol = 2, nrow = 1, which = 1:2) + theme_bw()





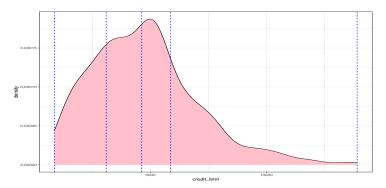
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mod <- lm(debt ~ credit_limit + income, data = credit_paradox)
summary(mod)\$coef</pre>

EstimateStd. Errort valuePr(>|t|)(Intercept)-385.179260419.464801525-19.788503.878764e-61credit_limit0.26432160.00587972944.954717.717386e-158income-7.66332300.385072058-19.901011.260933e-61

ID	debt	credit_limit	credit_rating	income	age	credit_cats
1	333	3606	283	14.891	34	(3.09e+03,4.62e+03]
2	903	6645	483	106.025	82	(5.87e+03,1.39e+04]
3	580	7075	514	104.593	71	(5.87e+03,1.39e+04]
4	964	9504	681	148.924	36	(5.87e+03,1.39e+04]
5	331	4897	357	55.882	68	(4.62e+03,5.87e+03]
6	1151	8047	569	80.180	77	(5.87e+03,1.39e+04]

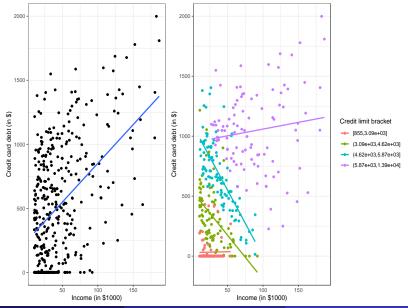
theme_bw()



```
credit_paradox %>%
  group_by(credit_cats) %>%
  summarize(n())
```

#	A tibble: 4 x 2	
	credit_cats	`n()`
	<fct></fct>	<int></int>
1	[855,3.09e+03]	100
2	(3.09e+03,4.62e+03]	100
3	(4.62e+03,5.87e+03]	100
4	(5.87e+03,1.39e+04]	100

```
p1 <- ggplot(data = credit_paradox, aes(x = income, y = debt))
  geom point() +
  geom smooth(method = "lm", se = FALSE) +
  theme bw() +
  labs(y = "Credit card debt (in $)",
       x = "Income (in $1000)")
p2 \leftarrow ggplot(data = credit paradox, aes(x = income, y = debt,
                                         color = credit cats))
  geom_point() +
  geom_smooth(method = "lm", se = FALSE) +
  theme bw() +
  labs(y = "Credit card debt (in $)",
       x = "Income (in $1000)",
       color = "Credit limit bracket")
p1 + p2
```



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